

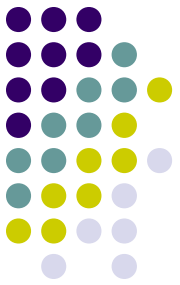
Forecasting Seasonal and Regional Typhoon Activity: A Track Based Approach

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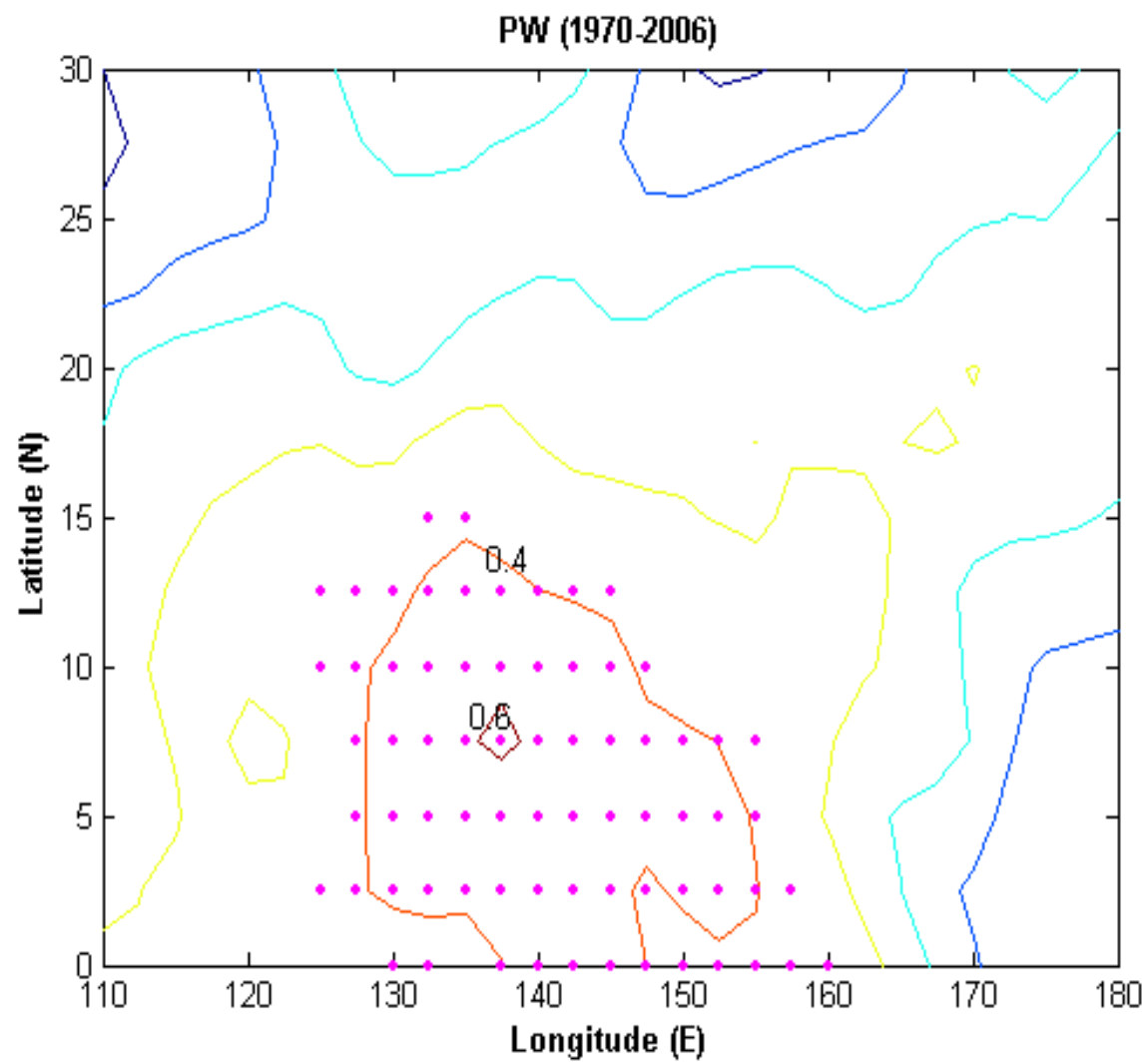
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- M.-M. Lu (Central Weather Bureau, Taiwan)
 - Part I: Probabilistic Forecasting (Bayesian)
 - Part II: Track-type Based Probabilistic Forecasting

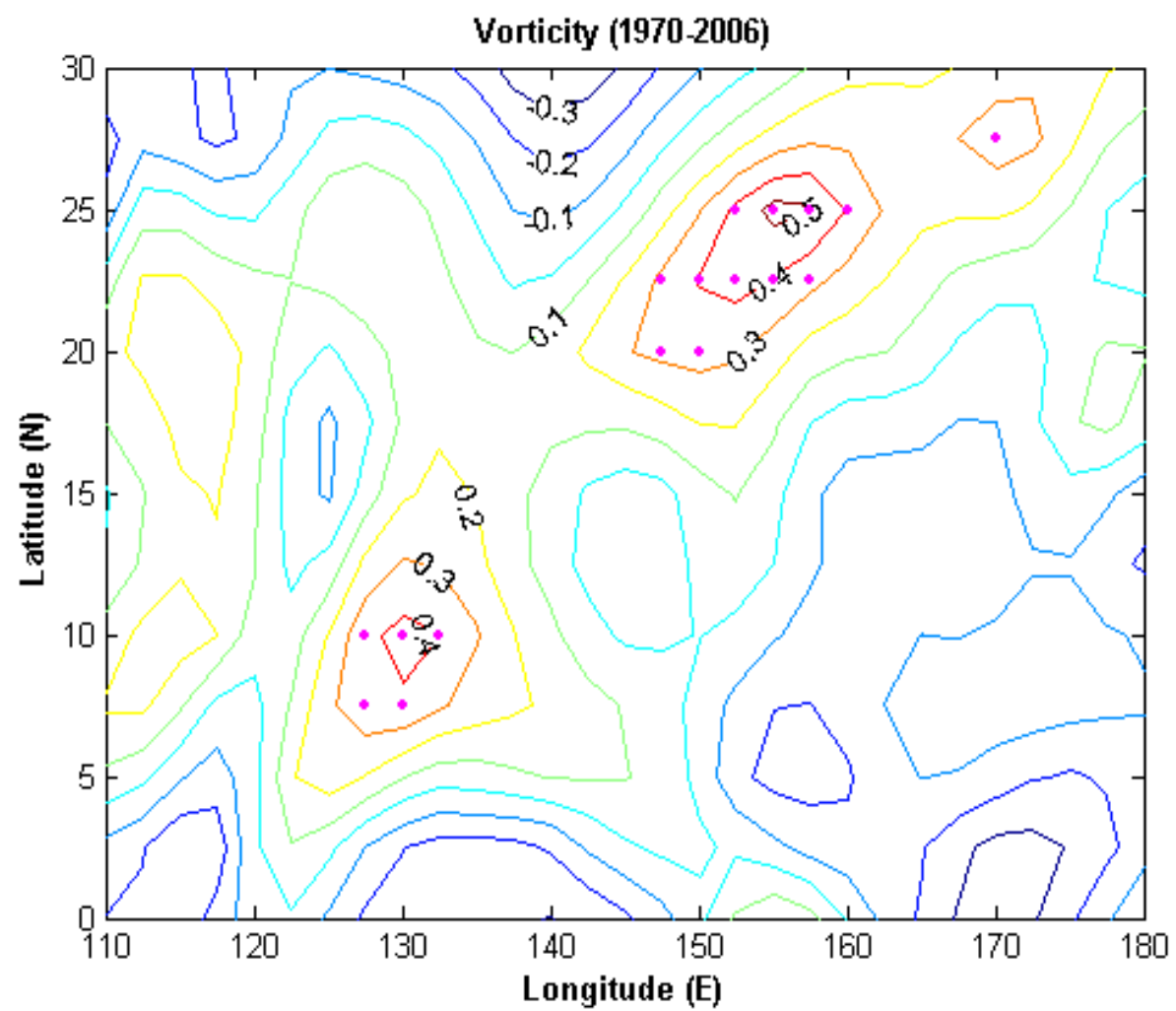
- Basin-wide (Chan et al. 1998 & 2001, Wea. Forecasting) to **regional** forecasts (Chu et al., 2007, TAO) for adequate planning of regional emergency management and hazard mitigation
- Deterministic (Chan et al., 1998, 2001; Chu et al., 2007) to **probabilistic** (Chu and Zhao, 2007, J. Climate) to facilitate decision-making

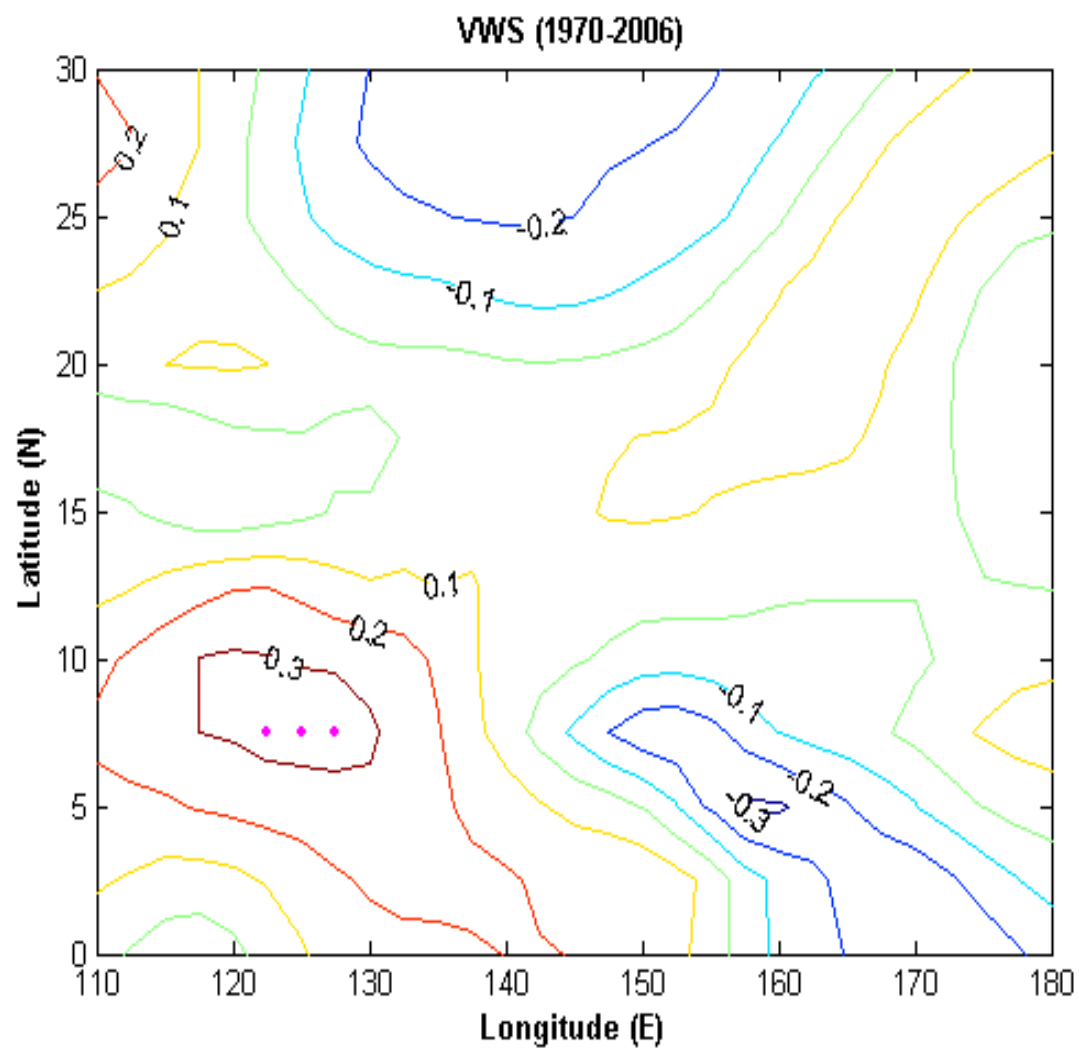


- Tropical cyclone data in the vicinity of Taiwan (21-26°N, 119-125°E) from 1970 to 2006 compiled by the CWB in Taiwan
- Monthly mean SST, wind data at 850- and 200-hPa levels, relative vorticity at the 850 hPa level, and total precipitable water over the tropical western North Pacific (NCEP/NCAR Reanalysis products)



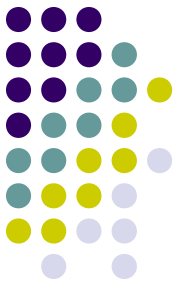






Part I: Probabilistic Forecasting

Poisson Distribution



Given the Poisson intensity parameter λ (i.e., the mean seasonal TC rates), the probability

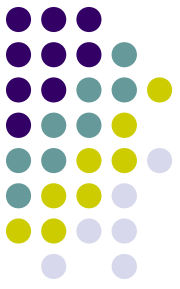
mass function (PMF) of h TCs occurring in T years is (Epstein, 1985)

$$P(h|\lambda, T) = \exp(-\lambda T) \frac{(\lambda T)^h}{h!} \quad (1)$$

$$h = 0, 1, 2, \dots \quad \lambda > 0 \quad T > 0$$

where λ and T

~~It may also be expressed in terms of the seasonal TC rates (Chan Yan, 2001; Hurrell, 2002). The TC distribution, based on the observed data, is given by the following equation:~~



Telesingapore

Telesingapore $i \in \mathbb{N}$, $Z_i \neq \emptyset$, Z_i is
the i -th element of the sequence Z .
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In the vector form, this model can be formulated as below:

$$P(h|Z) = \prod_{i=1}^N P(h_i|Z_i), \text{ where } h_i|Z_i \sim \text{Pois}(\beta_0 Z_i)$$

$$Z_i \sim N(\beta_1 + \beta_2 SST_i + \beta_3 PW_i + \beta_4 VOR_i + \beta_5 VWS_i, \sigma^2), \text{ where specifically}$$

$$X_i = [1, SST_i, PW_i, VOR_i, VWS_i], i = 1, 2, \dots, N$$

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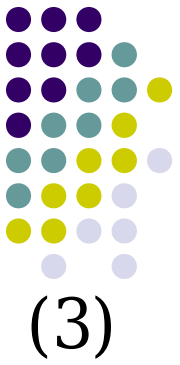
$$\beta = [\beta_0, \beta_1, \beta_2, \beta_3, \beta_4, \beta_5]'$$

Based on Bayes' theorem, it is obvious that

$$\begin{aligned} P(\mathbf{Z} | \mathbf{h}, \boldsymbol{\beta}, \sigma^2) &\propto P(\mathbf{h} | \mathbf{Z}, \boldsymbol{\beta}, \sigma^2) P(\mathbf{Z} | \boldsymbol{\beta}, \sigma^2) \\ &= P(\mathbf{h} | \mathbf{Z}) P(\mathbf{Z} | \boldsymbol{\beta}, \sigma^2) \end{aligned} \quad (3)$$

Substituting the probability model (2) into (3), ignoring the constant part, yields the conditional posterior distribution for a Poisson model

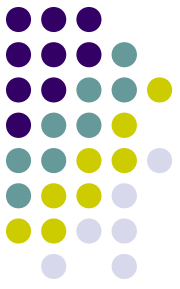
$$P(\mathbf{Z} | \mathbf{h}, \boldsymbol{\beta}, \sigma^2) \propto \frac{1}{\sigma^N} \prod_{i=1}^N \exp \left[-\left(e^{Z_i} + Z_i h_i \right) - \frac{1}{2\sigma^2} (Z_i - \mathbf{X}_i \boldsymbol{\beta})^2 \right] \quad (4)$$





Since we do not have any credible prior information for the coefficient vector $\boldsymbol{\beta}$ and the variance σ^2 , it is reasonable to choose the non-informative prior. In formula, it is (Gelman et al., 2004, p. 355)

$$P(\boldsymbol{\beta}, \sigma^2) \propto \sigma^{-2}$$



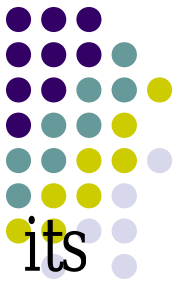
With the new observed predictor set $\tilde{\mathbf{X}} = [\tilde{X}_{i1}, \tilde{X}_{i2}, \dots, \tilde{X}_{iK}]$, if we have the posterior distribution of the parameters, the predictive distribution for the latent variable \tilde{Z} and IC counts \tilde{h} will be

$$P(\tilde{Z} | \tilde{X}, \mathbf{X}, \mathbf{h}) = \int \int_{\boldsymbol{\beta}, \sigma^2} P(\tilde{Z} | \tilde{X}, \boldsymbol{\beta}, \sigma^2) P(\boldsymbol{\beta}, \sigma^2 | \mathbf{X}, \mathbf{h}) d\boldsymbol{\beta} d\sigma^2$$

(5a)

$$P(\tilde{h} | \tilde{X}, \mathbf{X}, \mathbf{h}) = \int_{\tilde{Z}} \frac{\exp(-e^{\tilde{Z}} + \tilde{Z}\tilde{h})}{\tilde{h}!} P(\tilde{Z} | \tilde{X}, \mathbf{X}, \mathbf{h}) d\tilde{Z}$$

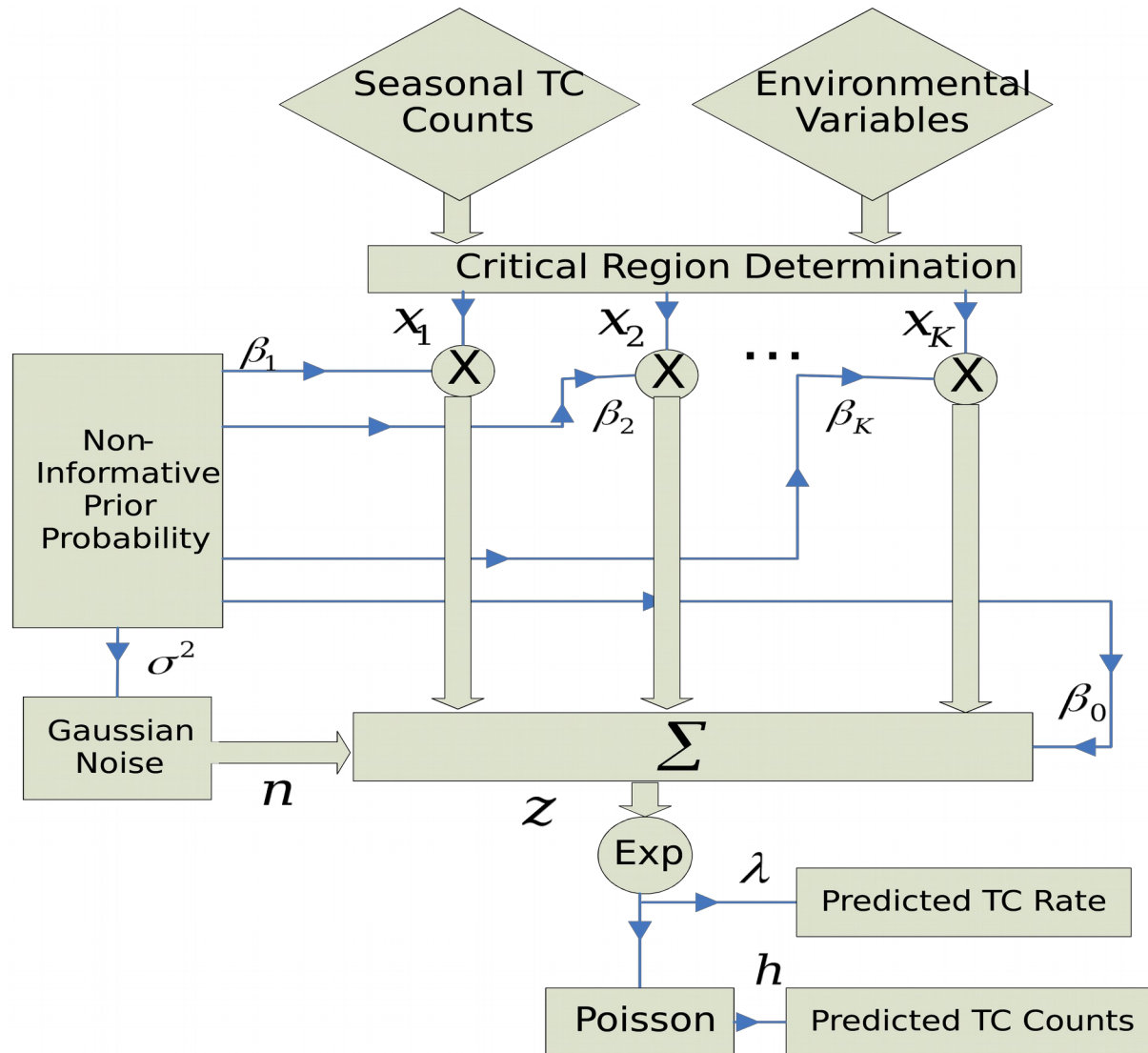
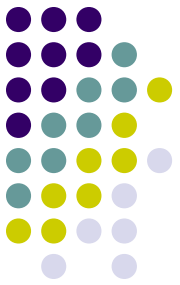
(5b)



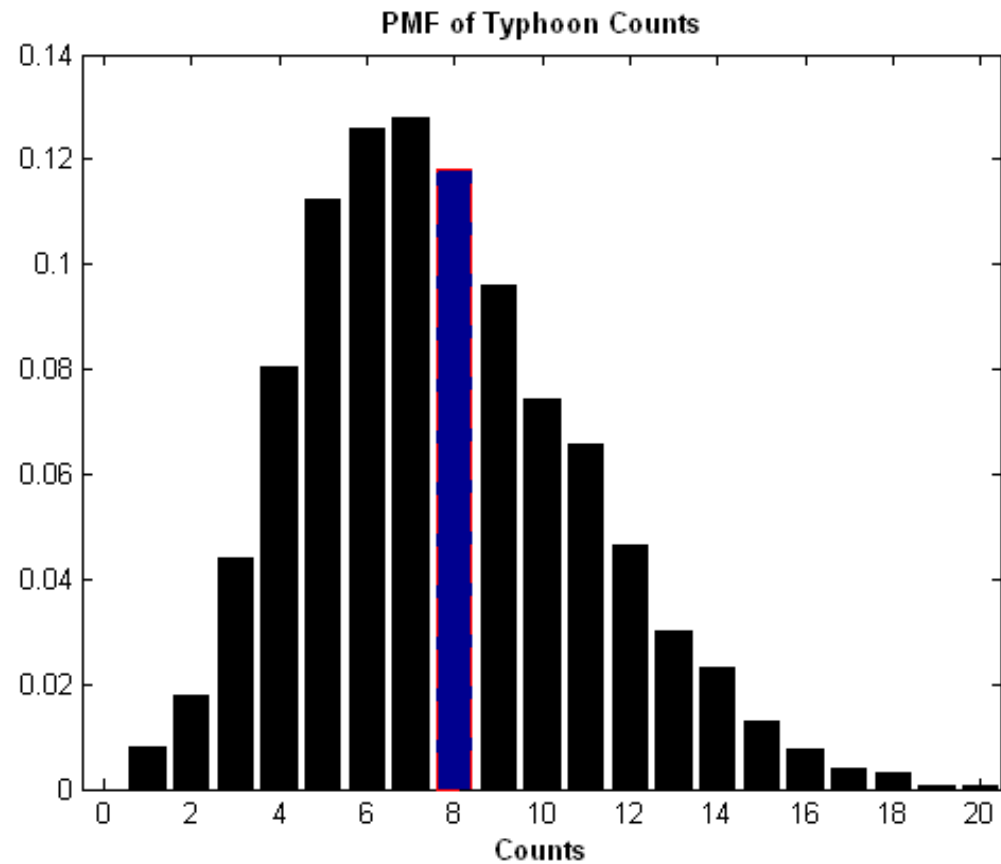
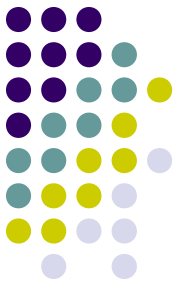
We design a Gibbs sampler, which has $P(\boldsymbol{\beta}, \sigma^2 | \mathbf{X}, \mathbf{h})$ as its stationary distribution, and then we can use an alternative approach to integrate (5a) by

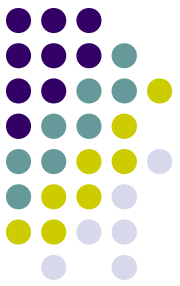
$$P(\tilde{Z} | \tilde{X}, \mathbf{X}, \mathbf{h}) = \frac{1}{L} \sum_{i=1}^L P(\tilde{Z} | \tilde{X}, (\boldsymbol{\beta}, \sigma^2)^{[i]}) \quad (6)$$

where $(\boldsymbol{\beta}, \sigma^2)^{[i]}$ is the i -th sampling from the Gibbs sampler after the burn-in period.



Forecasting a busy year (2004)

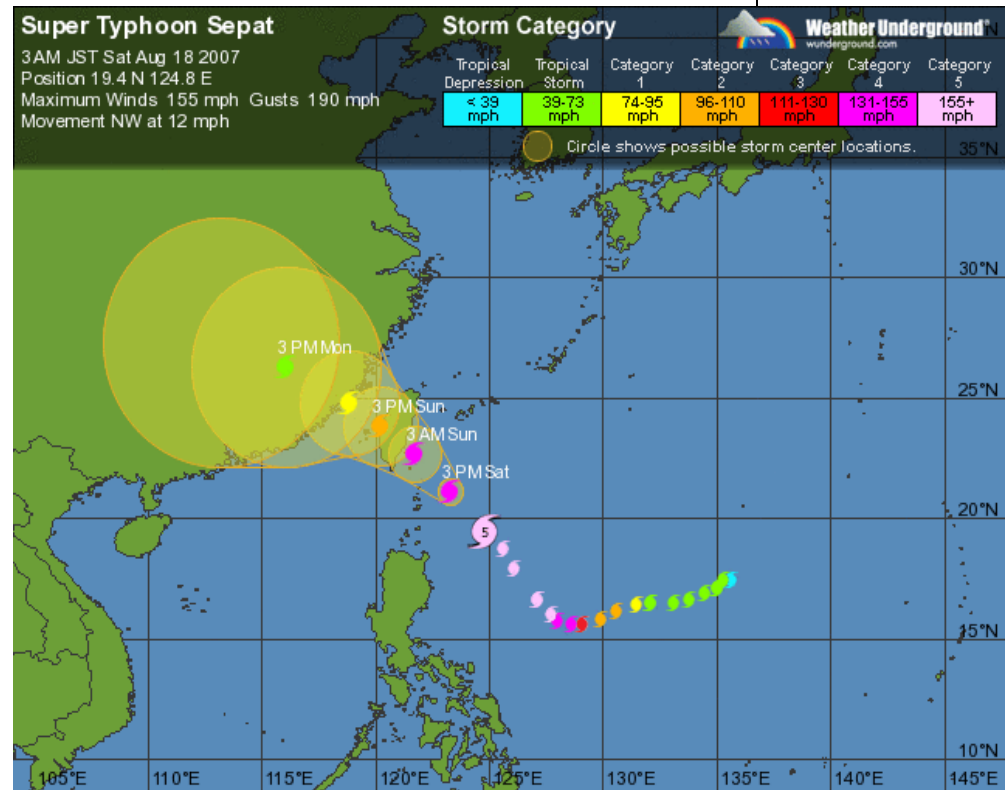


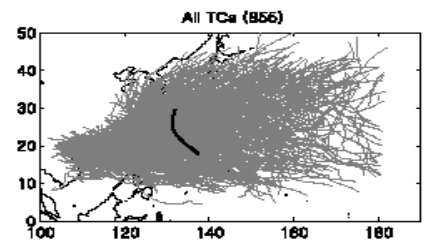
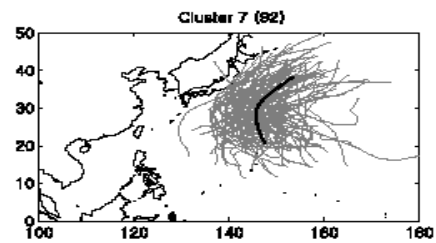
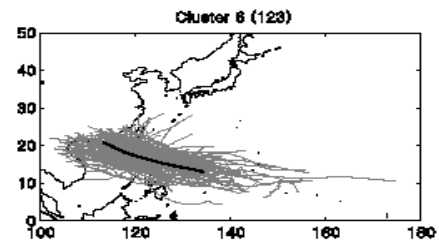
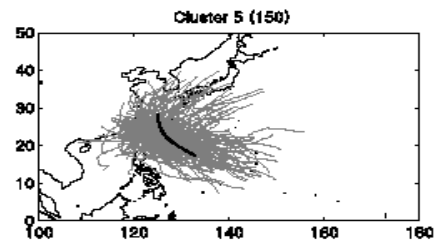
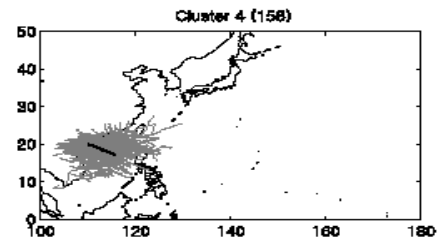
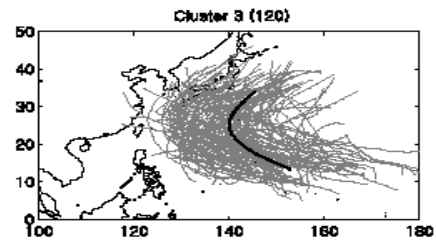
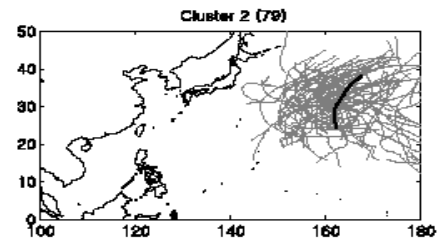
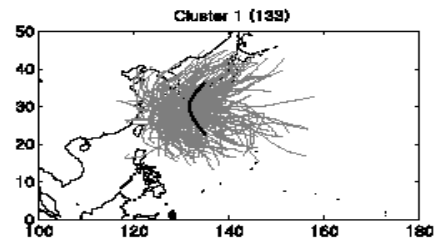


Part II: A Track-type based Approach

- A vector EOF analysis for TC tracks
- Fuzzy clustering of TC tracks (Harr and Elsberry, 1995)

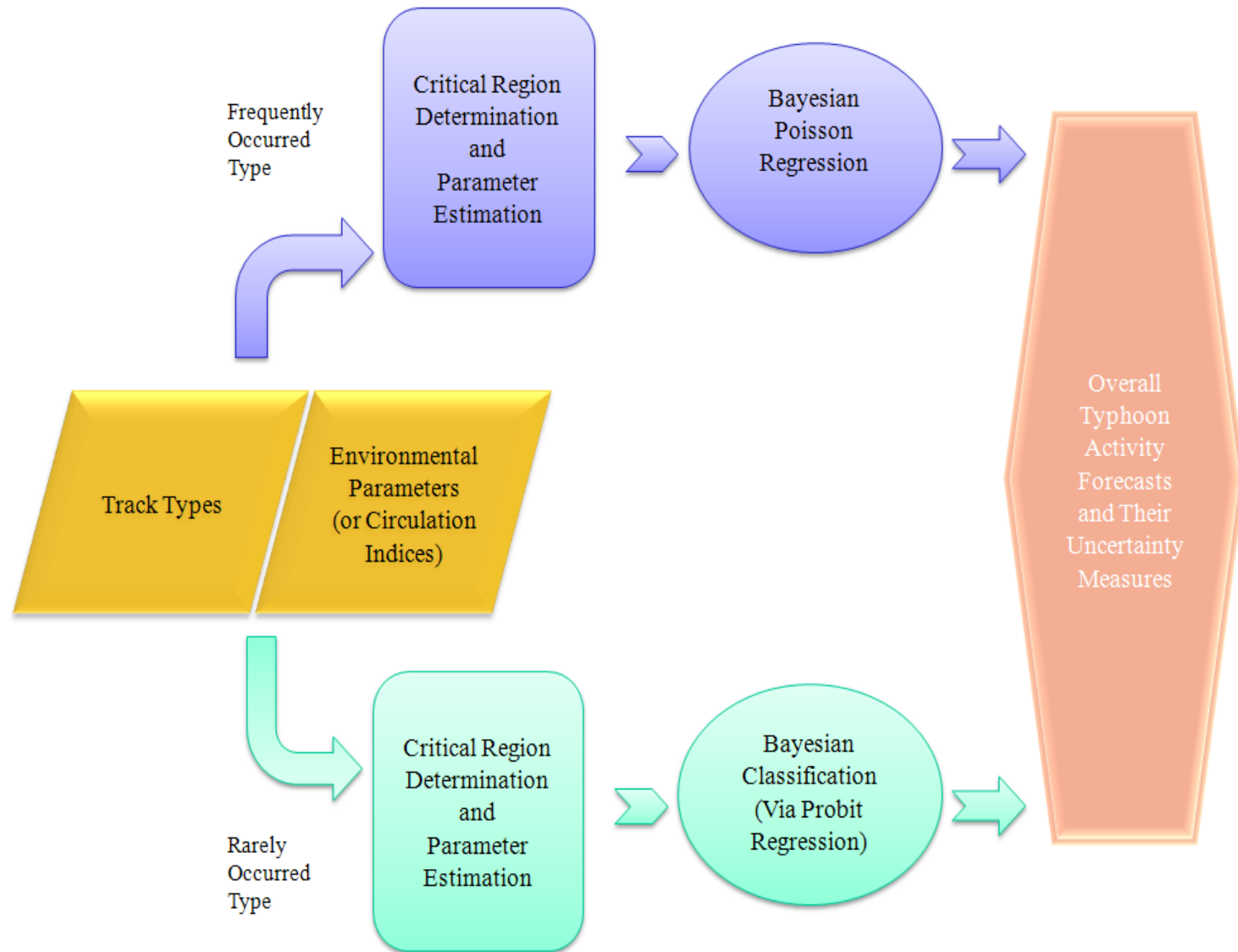
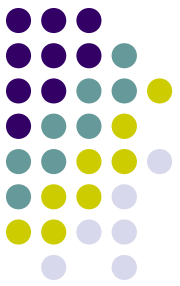
SNU group

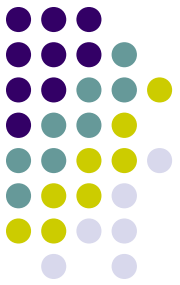




Year	Type 1	Type 2	Type 3	Type 4	Type 5	Type 6	Type 7	Total
1997	0	0	1	0	1	0	0	2
1998	1	0	0	1	2	0	0	4
1999	1	0	0	0	0	1	0	2
2000	1	0	0	0	4	1	0	6
2001	1	0	0	0	4	1	0	6
2002	0	0	1	0	3	0	0	4
2003	0	0	0	0	5	1	0	6
2004	0	0	0	1	6	1	0	8
2005	0	0	0	0	5	0	0	5
2006	0	0	0	0	4	1	0	5







Summary

- Two methods for predicting seasonal typhoon activity are introduced (Basin-wide to **regional**, deterministic to **probabilistic**).

A **state-of-the-art hierarchical Bayesian system** (i.e., Poisson or probit regression), rooted on the **track patterns**, is currently being developed to provide probabilistic forecasts for typhoons near Taiwan.

It may be possible to forecast mean genesis locations, mean path, and landfall locations for each track types.

